

PROBABILITY THEORY

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This is a *personal* geodesic along probability theory. It is not a textbook, nor an attempt at it. It does *not* constitute a syllabus for my course. Definitions, lemmas, and examples are often merged in the text, and it is up to the reader to discern which is which. Corrections, suggestions, observations, . . . , are most welcome. Version of December 19, 2024.

1. BASICS

Definition 1.1. Let $\emptyset \in \mathcal{C} \subseteq \mathcal{P}(\Omega)$, for some nonempty set Ω . Then \mathcal{C} is:

- (1) a *semialgebra* \mathcal{S} if it is closed under finite intersections (including the empty one, so $\Omega \in \mathcal{S}$) and, for every $A \in \mathcal{S}$, A^c is a finite disjoint union of elements of \mathcal{S} ;
- (2) an *algebra* \mathcal{A} if it is closed under all boolean operations;
- (3) a *σ -algebra* \mathcal{F} if it is closed under all boolean operations and countable unions.

Example 1.2. $\mathcal{P}(\Omega)$. Finite/cofinite subsets of an infinite set. Countable/cocountable subsets of \mathbb{R} . \emptyset, \mathbb{R} , plus all intervals $(-\infty, b], (a, +\infty), (a, b]$, for a, b in \mathbb{R} or in \mathbb{Q} .

If A, B belong to the semialgebra \mathcal{S} , then $A \setminus B$ is a finite disjoint union of elements of \mathcal{S} .

Note that an algebra is an algebra $(\mathcal{A}, \Delta, \cap, \emptyset, \Omega)$ in the algebraic sense of the word, over the field $F_2 = \{0, 1\} = \{\emptyset, \Omega\}$.

The pullback of any semialgebra/algebra/ σ -algebra by any function is a semialgebra/algebra/ σ -algebra. In particular, this holds for $\Omega' \subseteq \Omega$.

Lemma 1.3. *Given any nonempty $\mathcal{C} \subseteq \mathcal{P}(\Omega)$, the set of all finite intersections of elements of \mathcal{C} and complements of elements of \mathcal{C} is a semialgebra (in general larger than \mathcal{C} , even though \mathcal{C} might already be a semialgebra).*

Proof. Indeed, it contains \emptyset and Ω , and is closed under finite intersections. Also, for $n \geq 2$, $(A_1 \cap \dots \cap A_n)^c$ is the disjoint union of the $2^n - 1$ sets $A_1^{f(1)} \cap \dots \cap A_n^{f(n)}$, for $f : \{1, \dots, n\} \rightarrow \{\text{nothing}, c\}$ not identically nothing. \square

Lemma 1.4. *Let $(\Omega_i, \mathcal{S}_i)_{i \in I}$ be a family of semialgebras. Then the set of all finite intersections of $\pi_i^{-1}S_i$, for $i \in I$ and $S_i \in \mathcal{S}_i$, is a semialgebra in $\prod_{i \in I} \Omega_i$, called the semialgebra of cylinders.*

Proof. Look at $(\pi_1^{-1}[S_1] \cap \dots \cap \pi_t^{-1}[S_t])^c$. By the proof of Lemma 1.3, it is a finite disjoint union of intersections of stuff of the form $\pi_i^{-1}[S_i]$ or $(\pi_j^{-1}[S_j])^c$. We now observe that $\dots \cap (\pi_j^{-1}[S_j])^c \cap \dots = \dots \cap \pi_j^{-1}[\dot{\cup}_m S_{j,m}] \cap \dots = \dot{\cup}_m \dots \cap \pi_j^{-1}[S_{j,m}] \cap \dots$. \square

Example 1.5. The following are semialgebras.

- (1) *Cylinders* in \mathbb{R}^d : finite intersections of π_i^{-1} of $\emptyset, \mathbb{R}, (-\infty, b], (a, +\infty), (a, b]$ in \mathbb{R} .
- (2) *Cylinders* in $m^{\mathbb{Z}_{\geq 0}}$: finite intersections of π_i^{-1} of subsets of m .

(3) *Blocks* in $m^{\mathbb{Z}_{\geq 0}}$: the emptyset, plus all $[a_0, \dots, a_{t-1}]$'s, for $a_0, \dots, a_{t-1} \in m$.

Definition 1.6. The intersection of any nonempty family of algebras/ σ -algebras is an algebra/ σ -algebra; hence we may speak about the algebra/ σ -algebra $\mathcal{A}(\mathcal{C})/\mathcal{F}(\mathcal{C})$ *generated* by any \mathcal{C} . The Borel σ -algebra \mathcal{B} .

Example 1.7. The intersection of the semialgebras $\{\emptyset, \{a, b, c\}, \{a\}, \{b\}, \{c\}\}$ and $\{\emptyset, \{a, b, c\}, \{a\}, \{b, c\}\}$ is not a semialgebra.

Definition 1.8. of measurable map $F : (\Omega, \mathcal{F}) \rightarrow (\Omega', \mathcal{F}')$.

Lemma 1.9. Let $F : (\Omega, \mathcal{F}) \rightarrow (\Omega', \mathcal{F}')$ be a map, and let \mathcal{C}' be a subset of \mathcal{F}' that generates \mathcal{F}' as a σ -algebra. Assume that the F -counterimage of every element of \mathcal{C}' is in \mathcal{F} . Then F is measurable. In particular, a continuous map between topological spaces is a measurable map w.r.t. the relative Borel σ -algebras.

Proof. Just note that $\{A \subseteq \Omega' : F^{-1}A \in \mathcal{F}\}$ is a σ -algebra. \square

Theorem 1.10. The Borel σ -algebra of \mathbb{R} is generated both by the family of all open intervals with rational endpoints, and by the family of all intervals $(-\infty, a]$, with rational a 's. Hence, we can check measurability of functions just by checking the counterimages of $(-\infty, a]$'s.

Proof. Key point: $\mathcal{C} \subseteq \mathcal{F}(\mathcal{D})$ implies $\mathcal{F}(\mathcal{C}) \subseteq \mathcal{F}(\mathcal{D})$. We have $\{\text{open sets}\} \subseteq \mathcal{F}(\{\text{open intervals}\})$, and hence the first statement. For the second one, we need $\{(-\infty, a]\} \subseteq \mathcal{F}(\{\text{open intervals}\})$ and $\{\text{open intervals}\} \subseteq \mathcal{F}(\{(-\infty, a]\})$, which is easy. \square

Example 1.11. Both cylinders and blocks generate the Borel σ -algebra of $m^{\mathbb{Z}_{\geq 0}}$.

Lemma 1.12. The algebra $\mathcal{A}(\mathcal{S})$ generated by the semialgebra \mathcal{S} is the set \mathcal{Q} of all finite disjoint unions of elements of \mathcal{S} .

Proof. One direction is clear. For the other, it is enough to check that \mathcal{Q} is an algebra. Closure by finite intersections is clear. For complements, we have

$$(S_1 \dot{\cup} \dots \dot{\cup} S_t)^c = S_1^c \cap \dots \cap S_t^c = (T_{1,1} \dot{\cup} \dots \dot{\cup} T_{1,r_1}) \cap \dots \cap (T_{1,t} \dot{\cup} \dots \dot{\cup} T_{t,r_t}),$$

which is in \mathcal{Q} , since it is a finite intersection of elements of \mathcal{Q} . \square

Example 1.13. The algebra \mathcal{A} generated by blocks in $m^{\mathbb{Z}_{\geq 0}}$ equals the algebra generated by cylinders. We have $A \in \mathcal{A}$ iff A is clopen in $m^{\mathbb{Z}_{\geq 0}}$ iff A is expressible as a finite boolean combination of sets of the form $(\omega_i = a)$. For $m = 2$, \mathcal{A} is the free boolean algebra on countably many generators.

Definition 1.14. Let \mathcal{C} be either a semialgebra, or an algebra, or a σ -algebra on Ω . A [positive] *measure* on (Ω, \mathcal{C}) is a map $\mu : \mathcal{C} \rightarrow [0, +\infty]$ satisfying $\mu(\emptyset) = 0$ and conditional σ -additivity. If $\mu(\Omega) = 0$ then μ is *trivial*; if not otherwise specified, “measure” will always mean “nontrivial measure”. Relaxing σ -additivity to finite additivity we get the definition of a *finitely additive measure*. A measure is *finite* if $\mu(\Omega) < +\infty$, and is a *probability* if $\mu(\Omega) = 1$; in this case we use P for μ . A measure is σ -*finite* if Ω can be written as a countable union of μ -finite sets.

Example 1.15. The counting measure $(\Omega, \mathcal{P}(\Omega), \sharp)$ on a finite, countable, and uncountable set. Dirac measures. Measures are closed under finite or countable nonnegative combinations, and probabilities under finite or countable affine combinations.

Non-principal ultrafilters provide examples of $\{0, 1\}$ -valued finitely additive probabilities. Also, setting μ on a countable set to be $0/\infty$ on finite/infinite subsets we obtain a finitely additive measure which is not a measure.

Let μ be finitely additive on an algebra. Then $A \subseteq B$ implies $\mu(A) \leq \mu(B)$; if $\mu(B) < \infty$, then $\mu(B \setminus A) = \mu(B) - \mu(A)$. Moreover, $\mu(A \cup B) \leq \mu(A) + \mu(B)$ and $\mu(A \Delta B) = 0$ implies $\mu(A) = \mu(B) = \mu(A \cap B)$.

For every at most countable family $\{a_i\}_{i \in I}$ in $[0, +\infty]$, the sum $\sum_i a_i$ is well defined and does not depend on the order.

If Ω is at most countable, we always endow it with the σ -algebra $\mathcal{P}(\Omega)$; it is then easy to describe all measures and all probabilities on Ω .

Definition 1.16. Let μ be a f.a. measure on a semialgebra. If for every event and countable family of events $A, \{A_i\}_{i < \omega}$ we have

$$A = \bigcup_{i < \omega} A_i \text{ implies } \mu(A) \leq \sum_{i < \omega} \mu(A_i),$$

then we say that μ is σ -subadditive.

Theorem 1.17. Let μ be a f.a. measure on the algebra \mathcal{A} . T.f.a.e.:

- (1) μ is a measure;
- (2) for every $A, A_0, A_1, \dots \in \mathcal{A}$, if A_n converges monotonically increasing to A , then $\mu(A_n)$ converges to $\mu(A)$;
- (3) μ is σ -subadditive.

If the above conditions hold then:

- (4) for every $A, A_0, A_1, \dots \in \mathcal{A}$, if at least one of the A_n is μ -finite and A_n converges monotonically decreasing to A , then $\mu(A_n)$ converges to $\mu(A)$.

Proof. (1) implies (2): consider $B_i = A_i \setminus A_{i-1}$. (2) implies (3): consider $B_i = A_0 \cup \dots \cup A_i$ and assume $\sum_{i < \omega} \mu(A_i) = \alpha < \mu(A)$. Then all partial sums are $\leq \alpha$ and therefore $\mu(B_i)$ cannot converge to $\mu(A)$. (3) implies (1): let $A = \bigcup A_i$. We need $\mu(A) \geq \sum \mu(A_i)$, which is true since $\mu(A)$ is \geq than the measure of any finite union of the A_i 's.

(2) implies (4) by passing to the complements. \square

Given a sequence of events $(A_n)_{n < \omega}$, the functions $\limsup_n \mathbb{1}_{A_n}$ and $\liminf_n \mathbb{1}_{A_n}$ (defined componentwise) exist and are $\{0, 1\}$ -valued; hence they define sets

$$B = \limsup_n A_n = \bigcap_n \bigcup_{k \geq n} A_k,$$

$$C = \liminf_n A_n = \bigcup_n \bigcap_{k \geq n} A_k.$$

The function $\lim_n \mathbb{1}_{A_n}$ may or may not exist; if it exists it is $\{0, 1\}$ -valued and hence defines a set $D = \lim_n A_n$.

Lemma 1.18. B and C are events with $B \supseteq C$. $B = C$ iff D exists. If D exists, then it is an event. In that case, for every measure μ such that $\mu(\bigcup_{k \geq n} A_k)$ is finite for some n , we have $\mu(D) = \lim_n \mu(A_n)$.

Proof. For the last claim we have, for every n ,

$$\bigcap_{k \geq n} A_k \subseteq D \subseteq \bigcup_{k \geq n} A_k,$$

as well as

$$\bigcap_{k \geq n} A_k \subseteq A_n \subseteq \bigcup_{k \geq n} A_k.$$

□

2. INDEPENDENCE AND CONDITIONAL PROBABILITY

Definition 2.1. A μ -measurable partition of (X, \mathcal{X}, μ) is a finite or countable family $\{E_i\}_{i \in I}$ of elements of \mathcal{X} such that:

- (1) $E_i \cap E_j = \emptyset$ for $i \neq j$;
- (2) $\mu(X \setminus \bigcup E_i) = 0$.

We identify two μ -measurable partitions $\{E_i\}$ and $\{E'_i\}$ on the same set of indices if $\mu(E_i \triangle E'_i) = 0$ for every i .

For the rest of this section we deal with probability spaces (Ω, \mathcal{F}, P) only.

Definition 2.2. A family $\{A_i\}_{i \in I}$ of events is P -independent if for every finite $J \subseteq I$ we have

$$P\left(\bigcap_{j \in J} A_j\right) = \prod_{j \in J} P(A_j).$$

A family of pairwise P -independent elements is not necessarily P -independent. If A, B are P -independent, so are A, B^c .

Definition 2.3. Let $P(E) > 0$; we then have the probability space $(\Omega, \mathcal{F}, P(-|E))$, and $P(-|E)$ is the *conditional probability given E* .

The events A such that P and $P(-|E)$ agree on A are precisely those events which are P -independent with E .

Lemma 2.4. Fix A_1, \dots, A_n and assume $P(A_1 \cap \dots \cap A_{n-1}) > 0$. Then

$$\begin{aligned} P(A_1 \cap \dots \cap A_n) &= P(A_1 \cap \dots \cap A_{n-1}) \cdot P(A_n | A_1 \cap \dots \cap A_{n-1}) \\ &= P(A_1) \cdot P(A_2 | A_1) \cdot P(A_3 | A_1 \cap A_2) \cdots P(A_n | A_1 \cap \dots \cap A_{n-1}). \end{aligned}$$

We agree once for all that $\infty \cdot 0 = 0$. In particular, if $P(E) = 0$, then we have that $P(A|E)$ is undefined, but the product $P(A|E)P(E)$ is defined and equals 0.

Theorem 2.5 (The Bayes Theorem). Let $\{E_i\}_{i \in I}$ be a measurable partition.

(1)

$$P = \sum_i P(E_i)P(-|E_i).$$

(2) If $P(A) > 0$ then, for every i , we have

$$P(E_i | A) = P(A | E_i) \frac{P(E_i)}{P(A)}.$$

3. RANDOM VARIABLES

Lemma 3.1. Let $F : (\Omega, \mathcal{F}) \rightarrow (\Omega', \mathcal{F}')$ be a measurable map, and let μ be a measure on (Ω, \mathcal{F}) . Then $F_*\mu$ is a measure on (Ω', \mathcal{F}') .

Definition 3.2. A *random variable* is a measurable map $X : (\Omega, \mathcal{F}, P) \rightarrow (R, \mathcal{B})$, where R is any of the topological spaces $\mathbb{R}, \mathbb{R}^d, \mathbb{C}$. The probability X_*P on R is said to be the *distribution*, or the *law*, of X . The variable X is *discrete* if $X[\Omega]$ is discrete (and hence at most countable, since R has a countable basis) in R . It is *continuous* if $P(X = x) = (X_*P)(\{x\}) = 0$ for every $x \in R$.

Definition 3.3. Let μ be a measure on (R, \mathcal{B}) , for R one of $\mathbb{R}, \mathbb{R}^d, \mathbb{C}$. The *support* of μ is the complement of the set of all x such that x belongs to an open set of μ -measure 0. Of course $\text{supp}(\mu)$ is closed, and is a subset of the closure of $X[\Omega]$.

Definition 3.4. Given any Borel probability μ on \mathbb{R} , its *cumulative distribution function*, or *repartition function*, is the function $M : \mathbb{R} \rightarrow [0, 1]$ defined by $M(x) = \mu((-\infty, x])$. Note that $\mu((-\infty, x)) = \lim_{x' \nearrow x} M(x') = M(x - 0)$ and $\mu(\{x\}) = M(x) - M(x - 0)$.

If μ is discrete, then it has strictly positive value only in the points of the finite or countable set $A = \text{supp}(\mu)$. Its *discrete-density distribution function* is then the function $m : \mathbb{R} \rightarrow \mathbb{R}_{\geq 0}$ defined by

$$m(x) = \begin{cases} \mu(\{x\}), & \text{if } x \in A; \\ 0, & \text{otherwise.} \end{cases}$$

Example 3.5. The following are examples of discrete distributions on \mathbb{R} ; plots done with SageMath, <http://www.sagemath.org/>.

(1) The *Binomial*, or *Bernoulli* distribution, $\text{Bin}(n, p)$, for $n \in \mathbb{Z}_{>0}$ and $p \in (0, 1)$:

$$m(k) = \binom{n}{k} p^k (1-p)^{n-k}.$$

(2) The *Hypergeometric* distribution, $\text{Hyp}(N, K, n)$, for $1 \leq K, n \leq N$, see Figure 1:

$$m(k) = \frac{\binom{K}{k} \binom{N-K}{n-k}}{\binom{N}{n}}.$$

(3) The *Poisson* distribution, $\text{Poisson}(\mu)$, for $\mu \in \mathbb{R}_{>0}$, see Figure 2:

$$m(k) = \exp(-\mu) \frac{\mu^k}{k!}.$$

(4) The *Geometric* distribution, $\text{Geom}(p)$, for $p \in (0, 1)$:

$$m(k) = p(1-p)^k.$$

(5) The *Zeta* distribution, $\text{Zeta}(\alpha)$, for $\alpha \in \mathbb{R}_{>1}$:

$$m(k) = \zeta(\alpha)^{-1} \frac{1}{k^\alpha}.$$

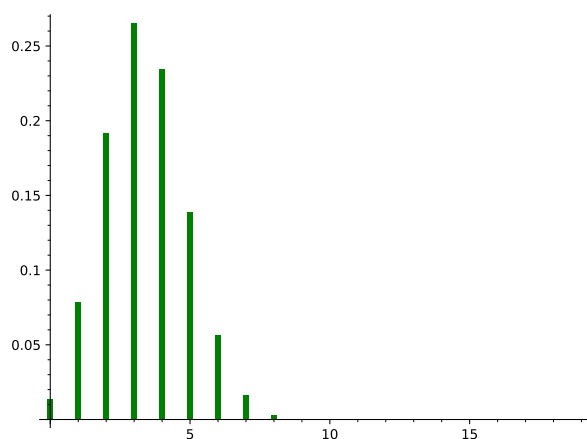


FIGURE 1

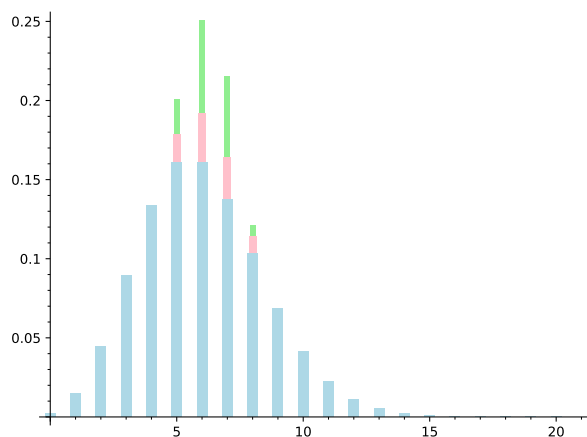


FIGURE 2

4. CONSTRUCTION OF MEASURES

Theorem 4.1 (The Monotone Class Theorem). *Let $\mathcal{C} \subseteq \mathcal{P}(\Omega)$ be closed under finite intersections (in particular, $\Omega \in \mathcal{C}$). Let \mathcal{M} be the smallest overclass of \mathcal{C} which is closed under countable increasing unions and increasing differences. Then $\mathcal{M} = \mathcal{F}(\mathcal{C})$.*

Proof. [JP03, Theorem 6.2 p. 36] □

Lemma 4.2. *Let \mathcal{A} be the algebra generated by the semialgebra \mathcal{S} , and let μ be a f.a. measure on \mathcal{S} . Then μ can be extended uniquely to a f.a. measure on \mathcal{A} . If μ is σ -subadditive on \mathcal{S} , then the extension remains σ -subadditive, and hence is a measure on \mathcal{A} .*

Theorem 4.3 (The Carathéodory-Hahn-Kolmogorov Extension Theorem). *Let \mathcal{F} be the σ -algebra generated by the algebra \mathcal{A} . Then every σ -finite measure on \mathcal{A} can be uniquely extended to \mathcal{F} , and the extension is again σ -finite.*

Proof. If an extension exists, then it is clearly σ -finite. Uniqueness is easy: let μ, ν be measures on \mathcal{F} that agree on \mathcal{A} . For every $A \in \mathcal{A}$ such that $\mu(A) = \nu(A)$ is finite, let $\mathcal{M}_A = \{B \in \mathcal{F} : \mu(B \cap A) = \nu(B \cap A)\}$. Then \mathcal{M}_A is a monotone overclass of \mathcal{A} , and hence equals \mathcal{F} . Using σ -finiteness we can write $\Omega = \bigcup_{i < \omega} A_i$ with $A_i \in \mathcal{A}$ of finite ($\mu = \nu$)-measure. Then, for every $B \in \mathcal{F}$, $\mu(B) = \sum_i \mu(B \cap A_i) = \sum_i \nu(B \cap A_i) = \nu(B)$. Existence is difficult: see [Bil95, Theorems 11.2 and 11.3]. \square

Lemma 4.4. *Let μ, ν be measures on (X, \mathcal{X}) , and let \mathcal{S} be a semialgebra that generates \mathcal{X} . Assume that $\mu = \nu$ on \mathcal{S} and that X is the union of at most countably many elements of \mathcal{S} of finite $\mu(= \nu)$ -measure. Then $\mu = \nu$.*

Corollary 4.5. *There exists a unique σ -finite measure on $(\mathbb{R}, \mathcal{B})$ that assigns to each interval its length. This measure is translation-invariant; it is named the Lebesgue measure, and denoted by λ .*

Proof. Let \mathcal{S} be the semialgebra of Example 1.2. Define λ to have value $a - b$ on $(b, a]$, value 0 on \emptyset , and $+\infty$ on all unbounded intervals. Then λ is a f.a., σ -subadditive measure on \mathcal{S} (additivity is clear, while σ -subadditivity requires a compactness argument). Hence it extends uniquely to a measure on the generated σ -algebra, namely \mathcal{B} . For translation-invariance, fix r and let $\mu = (T_r)_* \lambda$. Then $\mu = \lambda$ on \mathcal{S} . \square

Corollary 4.6. *Given any probability vector (p_0, \dots, p_{n-1}) , there exists a unique probability on (n^ω, \mathcal{B}) that assigns to each block $[a_0, \dots, a_{t-1}]$ the number $p_{a_0} \cdots p_{a_{t-1}}$.*

5. PROBABILITY MEASURES ON \mathbb{R}

A function $M : \mathbb{R} \rightarrow \mathbb{R}$ is *right continuous* in c if for every $\varepsilon > 0$ there exists $\delta > 0$ such that for every $x \in [c, c + \delta)$ we have $|Mx - Mc| < \varepsilon$. If M is nondecreasing, then M is right continuous in c iff for every sequence $x_0 \geq x_1 \geq \dots$ converging to c the sequence Mx_n converges to Mc iff there exists a sequence $x_0 \geq x_1 \geq \dots$ converging to c such that Mx_n converges to Mc ,

Probability measures on \mathbb{R} are completely described by the following theorem.

Theorem 5.1. *A function $M : \mathbb{R} \rightarrow [0, 1]$ is the repartition function of a probability μ on \mathbb{R} iff M is nondecreasing, right continuous, tending to 0 for $x \rightarrow -\infty$ and to 1 for $x \rightarrow +\infty$. If this happens, then μ and M determine each other.*

Remark 5.2. We have $\{b\} = \bigcap_{a \nearrow b} (a, b]$, and therefore $\mu(\{b\}) = \lim_{a \nearrow b} (Mb - Ma) = Mb - M(b - 0)$, with $M(b - 0) = \sup\{Mx : x < b\}$. Thus the random variable X is continuous according to Definition 3.2 iff the repartition function of X_*P is continuous.

Example 5.3. The following are examples of continuous probability distributions on \mathbb{R} which are induced by a Riemann-integrable density function.

- (1) The *Uniform* distribution with parameters $\alpha < \beta$, given by $m = (\beta - \alpha)^{-1} \mathbb{1}_{[\alpha, \beta]}$.
- (2) The *Exponential* distribution with parameter $\beta > 0$, given by $m(x) = \beta \exp(-\beta x)$ for $x > 0$, and $m(x) = 0$ otherwise.
- (3) More generally, the *Gamma* distribution with parameters $\alpha, \beta > 0$, given by

$$m(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} \exp(-\beta x)$$

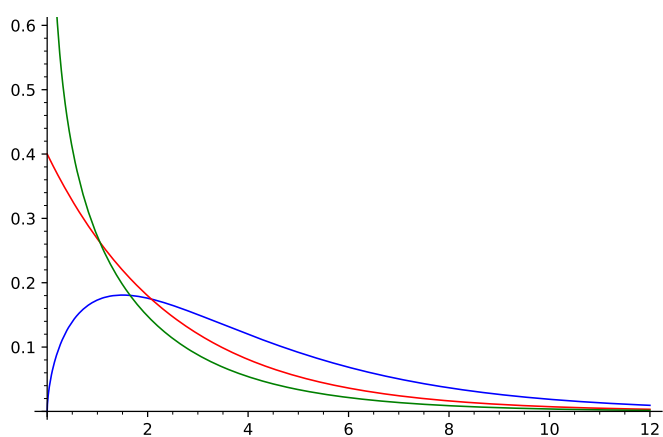


FIGURE 3

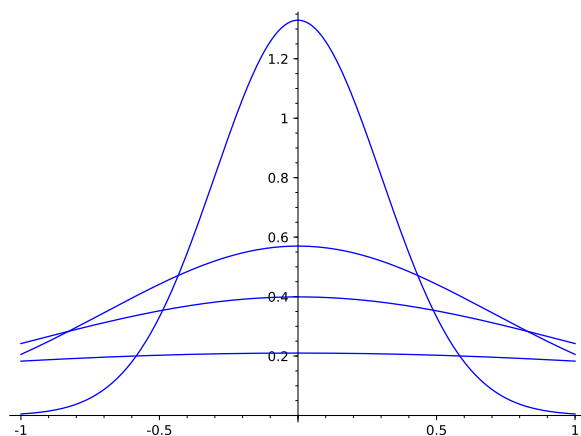


FIGURE 4

if $x > 0$, and $m(x) = 0$ otherwise; see Figure 3.

(4) The *Normal*, or *gaussian* distribution, $Normal(\mu, \sigma^2)$; see Figure 4.

$$m(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right).$$

Example 5.4. Repartition functions can be extremely slippery. Let P be the $(5/7, 2/7)$ -probability on $\Omega = 2^{\mathbb{N}}$, and let $X : \Omega \rightarrow \mathbb{R}$ be the random variable induced by the binary expansion. The repartition function of $\mu = X_*P$ is a typical devil's staircase; see Figure 5.

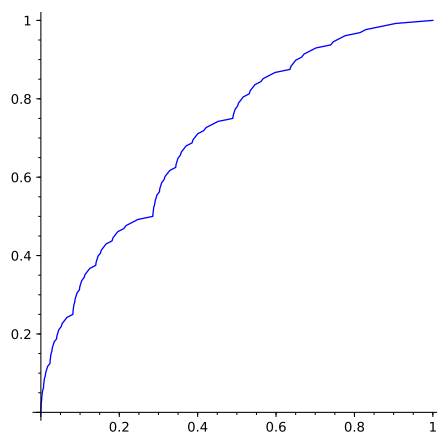


FIGURE 5

6. INTEGRATION THEORY

We see $[0, +\infty]$ and $[-\infty, +\infty]$ as topological spaces, homeomorphic to $[0, 1]$ and $[-1, 1]$ via, e.g., the cotangent function. They are then endowed with the Borel σ -algebra \mathcal{B} . We agree that $0 \cdot (+\infty) = 0$.

Lemma 6.1. *Let $f_0, f_1, \dots : (X, \mathcal{X}) \rightarrow ([-\infty, +\infty], \mathcal{B})$ be measurable.*

- (1) $\sup f_n, \inf f_n, \limsup f_n, \liminf f_n$ (all of them pointwise defined) are all measurable.
- (2) If the pointwise limit $\lim f_n$ exists, then it is measurable.
- (3) If $f_1, \dots, f_d : X \rightarrow \mathbb{R}$ are measurable, then $(f_1, \dots, f_d) : X \rightarrow \mathbb{R}^d$ is measurable.
- (4) If, moreover, $g : \mathbb{R}^d \rightarrow \mathbb{R}$ is measurable, then $g(f_1, \dots, f_d) : X \rightarrow \mathbb{R}$ is measurable; in particular, the set of measurable functions from X to \mathbb{R} is an \mathbb{R} -algebra.

Proof. $\{x : (\sup f_n)(x) \leq a\} = \bigcap_n \{x : f_n(x) \leq a\}$ and $\sup f_n$ is measurable. Analogously $\inf f_n$ is measurable. We have $\limsup_n f_n = \inf_n(\sup_{m \geq n} f_m)$, which is thus measurable. One proves the last two items by noting that the set of all products of open intervals with rational endpoints is a countable basis for the topology of \mathbb{R}^d , and therefore generates the Borel σ -algebra. \square

Definition 6.2. Let (X, \mathcal{X}, μ) be a σ -finite measure space. A *step function* is a measurable function $s : X \rightarrow \mathbb{R}_{\geq 0}$ whose range is finite; equivalently, it is a function that can be written (nonuniquely) as a finite sum $s = \sum_{i < n} a_i \mathbb{1}_{A_i}$.

Lemma 6.3. *Every step function can be written in disjoint form $s = \sum_{j < m} b_j \mathbb{1}_{B_j}$ in such a way that*

$$\sum_{i < n} a_i \mu(A_i) = \sum_{j < m} b_j \mu(B_j), \text{ (it may be } +\infty).$$

If $\sum_{j < m} b_j \mathbb{1}_{B_j} = \sum_{k < t} c_k \mathbb{1}_{C_k}$, then $\sum_{j < m} b_j \mu(B_j) = \sum_{k < t} c_k \mu(C_k)$.

Definition 6.4. Let $s = \sum_{i < n} a_i \mathbb{1}_{A_i} : X \rightarrow \mathbb{R}_{\geq 0}$ be a step function. We define

$$\mu(s) = \int_X s \, d\mu = \sum_{i < n} a_i \mu(A_i) \in [0, +\infty],$$

and note that $\mu(-)$ is positively linear.

Let $f : X \rightarrow [0, +\infty]$ be measurable. We define

$$\mu(f) = \int_X f \, d\mu = \sup\{\mu(s) : 0 \leq s \leq f\} \in [0, +\infty],$$

and note that $f \leq g$ implies $\mu(f) \leq \mu(g)$.

Lemma 6.5. Let $f : X \rightarrow [0, +\infty]$ be measurable.

- (1) There exists an increasing sequence $s_0, s_1, \dots : X \rightarrow \mathbb{R}_{\geq 0}$ of step functions that converges pointwise to f .
- (2) For every such sequence, $\mu(s_n)$ converges to $\mu(f)$ (this is a preliminary version of the Monotone Convergence Theorem).
- (3) $\mu(-)$ is positively linear (i.e., if g is another such function and $r \geq 0$, then $\mu(f + g) = \mu(f) + \mu(g)$ and $\mu(rf) = r\mu(f)$).

For $f : X \rightarrow [-\infty, +\infty]$ we set $f^+ = f \vee 0$, $f^- = (-f) \vee 0$, $|f| = f^+ + f^- = f^+ \vee f^- = f \vee (-f)$.

If at least one of $\mu(f^+)$, $\mu(f^-)$ is finite, then we say that f is *integrable* w.r.t. μ , and set

$$\int f \, d\mu = \int f(x) \, d\mu(x) = \int f(x) \mu(dx) = \int f^+ \, d\mu - \int f^- \, d\mu \in [-\infty, +\infty].$$

Since $f^+, f^- \leq |f| = f^+ + f^-$, we have that both of $\mu(f^+)$ and $\mu(f^-)$ are finite iff so is $\mu(|f|)$.

Definition 6.6. If $f = f_1 + if_2 : X \rightarrow \mathbb{C}$ is measurable, we say that $f \in \mathcal{L}_1(\mu) = \mathcal{L}_1(\mu, \mathbb{C})$ if $\mu(|f|) < +\infty$. Since $|f_1|, |f_2| \leq |f| \leq |f_1| + |f_2|$, this is equivalent to $f_1, f_2 \in \mathcal{L}_1(\mu, \mathbb{R})$; we then set $\mu(f) = \mu(f_1) + i\mu(f_2)$.

Theorem 6.7. (1) \mathcal{L}_1 is a complex vector space, and $\mu : \mathcal{L}_1 \rightarrow \mathbb{C}$ is a positive \mathbb{C} -linear functional.

- (2) If $f \in \mathcal{L}_1$, then $|\mu(f)| \leq \mu(|f|)$.

Theorem 6.8. Let $f : X \rightarrow \mathbb{C}$ be a measurable function with finite or countable range. Then $f \in \mathcal{L}_1$ iff the series

$$\sum \{a\mu(f^{-1}a) : a \in \text{range}(f)\}$$

is absolutely convergent (this clearly requires that $\mu(f^{-1}a) < +\infty$ for every $a \neq 0$). If this happens, then $\mu(f)$ equals the sum of that series.

Theorem 6.9. Let $R : (X, \mathcal{X}, \mu) \rightarrow (Y, \mathcal{Y})$ and $f : (Y, \mathcal{Y}) \rightarrow (\mathbb{C}, \mathcal{B})$ be measurable. Then $f \in \mathcal{L}_1(R_*\mu)$ iff $f \circ R \in \mathcal{L}_1(\mu)$. If this happens (or if f is $[0, +\infty]$ -valued), then

$$\int_X f \circ R \, d\mu = \int_Y f \, d(R_*\mu).$$

Proof. Case 1: f is a step function. Case 2: f is $[0, +\infty]$ -valued. Use the MCT for step functions. Case 3: \mathbb{R} -valued. Case 4: \mathbb{C} -valued. \square

Theorem 6.10 (The Monotone Convergence Theorem). *Let $f_0 \leq f_1 \leq \dots$ be a monotone sequence of $\mathbb{R}_{\geq 0}$ -valued measurable functions, and let $f = \lim f_n$. Then $\mu(f) = \lim \mu(f_n)$.*

Theorem 6.11 (The Fatou Lemma). *Let f_0, f_1, \dots be a sequence of $\mathbb{R}_{\geq 0}$ -valued measurable functions. Then $\mu(\liminf f_n) \leq \liminf \mu(f_n)$.*

Theorem 6.12 (The Dominated Convergence Theorem). *Let f_0, f_1, \dots be a sequence of \mathbb{C} -valued measurable functions. Assume that $\lim f_n = f$ (μ -a.e.) exists, and that there exists an $\mathbb{R}_{\geq 0}$ -valued $g \in \mathcal{L}_1$ such that $|f_n| \leq g$ for every n . Then all f_n s and f are in \mathcal{L}_1 , and $\lim \mu(f_n) = \mu(f)$.*

Theorem 6.13 (The Markov and Chebishev Inequalities). *Let $f : (X, \mu) \rightarrow \mathbb{R}_{\geq 0}$ be measurable, $a \geq 0$. Then:*

- (1) $a\mu(f \geq a) \leq \mu(f)$;
- (2) for every $p \geq 1$, we have $a^p\mu(f \geq a) \leq \mu(f^p)$.

Proof. Observe that $a\mathbb{1}_{(f \geq a)} \leq f$ and integrate. We obtain (2) by applying (1) to f^p and a^p . \square

Theorem 6.14. *Let $f_0, f_1, \dots : X \rightarrow \mathbb{C}$ be measurable.*

- (1) *If they are $\mathbb{R}_{\geq 0}$ -valued, then $\sum_n f_n$ is a $[0, +\infty]$ -valued measurable function, and*

$$\begin{aligned} \int_X \left(\int_{\mathbb{Z}_{\geq 0}} f_n(x) d\sharp(n) \right) d\mu(x) &= \int_X \left(\sum_n f_n \right) d\mu \\ &= \sum_n \int_X f_n d\mu = \int_{\mathbb{Z}_{\geq 0}} \left(\int_X f_n(x) d\mu(x) \right) d\sharp(n). \end{aligned} \quad (6.1)$$

- (2) *Assume that $\mu(\sum_n |f_n|) = \sum_n \mu(|f_n|) < +\infty$. Then:*
 - (i) $\sum_n f_n$ converges absolutely a.e., and determines a function in $\mathcal{L}_1(\mu)$.
 - (ii) (6.1) holds (this is a preliminary form of the Fubini theorem), and the series to the right converges absolutely;
 - (iii) $\lim_n f_n = 0$ a.e..

7. L_p SPACES

In this section $f, g, \dots : (X, \mathcal{X}, \mu) \rightarrow (\mathbb{C}, \mathcal{B})$ are complex-valued measurable functions. Let $1 \leq p < +\infty$. We set $\mathcal{L}_p(\mu) = \{f : f^p \in \mathcal{L}_1(\mu)\}$. Let $\|f\|_p = (\mu(|f|^p))^{1/p}$. By Minkowski's inequality $\|f + g\|_p \leq \|f\|_p + \|g\|_p$; in particular, $\mathcal{L}_p(\mu)$ is a \mathbb{C} -vector space. Moreover, $\|\cdot\|_p$ is a seminorm on $\mathcal{L}_p(\mu)$, and a norm on the set $L_p(\mu)$ of equivalence classes.

Lemma 7.1. *Let $f, g \in \mathcal{L}_p$. Then $\|f - g\|_p = 0$ iff $\mu(f \neq g) = 0$.*

Proof. $\mu(f \neq g) = 0$ implies $|f - g|^p = 0$ (a.e.) and $\int |f - g|^p = 0$. Conversely, for every $n \geq 1$, by Chebishev $(1/n)^p \mu(|f - g| \geq 1/n) \leq \int |f - g|^p$. Therefore $\mu(|f - g| \geq 1/n) = 0$ for every n , and $\mu(f \neq g) = 0$. \square

Let $\|f\|_\infty$ be the infimum of all $M \in [0, +\infty]$ such that $\mu(M < |f|) = 0$; sometimes $\|f\|_\infty$ is called the *essential supremum* of f . Let $\mathcal{L}_\infty(\mu) = \{f : \|f\|_\infty < +\infty\}$. Again $\|\cdot\|_\infty$ is a seminorm, and again $\|f - g\|_\infty = 0$ amounts to $\mu(f \neq g) = 0$. Thus the set $L_\infty(\mu)$ of equivalence classes inherits from $\mathcal{L}_\infty(\mu)$ a complex vector space structure.

All \mathcal{L}_p 's and L_p 's are closed under complex conjugation.

Theorem 7.2. *Let $f, g \in L_2(\mu)$. Then $fg \in L_1(\mu)$, and the Cauchy-Schwarz-Bunyakovsky-Hölder inequality*

$$|\langle f, g \rangle| := \left| \int_X \bar{f}g \, d\mu \right| \leq \|f\|_2 \|g\|_2$$

holds.

Theorem 7.3. *Let P be a probability; then*

$$L_\infty(P) \subseteq \cdots \subseteq L_3(P) \subseteq L_2(P) \subseteq L_1(P).$$

8. $E(X)$, $\text{Var}(X)$, $G_X(z)$ FOR COMMON R.V.S

Suppose $X \in L_p(P)$, with $p < \infty$. Then $E(X^p)$ is the p th moment of X , and $E((X - E(X))^p)$ its p th central moment.

If $X \in L_2(P)$ and $E(X) = \mu$ (= expectation = mean), define

$$\text{Var}(X) = E(|X - \mu|^2) = \sigma^2 \in \mathbb{R}_{\geq 0}.$$

We have:

- (1) $\text{Var}(X) = E(|X|^2) - |\mu|^2$.
- (2) For $a \geq 0$, $a^2 P(|X - \mu| \geq a) \leq \sigma^2$.
- (3) $\text{Var}(X + a) = \text{Var}(X)$ and $\text{Var}(aX) = |a|^2 \text{Var}(X)$.
- (4) If $\text{Var}(X) = 0$, then $X = \mu \mathbb{1}$ (a.e.).

Definition 8.1. Let X have values in $\mathbb{Z}_{\geq 0}$, with $X_*P = \mu$ and discrete-density function m . Then, for every z in the complex closed unit disk, the series

$$G_X(z) = \sum_{k=0}^{\infty} m(k)z^k = \int_{\mathbb{R}} z^x \, d\mu(x) = \int_{\Omega} z^{X(\omega)} \, dP(\omega) = E(z^X)$$

(written also G_μ or G_m) converges absolutely, and determines the *generating function* of X .

Since G can be differentiated termwise inside its disk of convergence, it determines μ via

$$m(k) = \frac{G^{(k)}(0)}{k!}.$$

We have $G(1) = 1$ and, if the radius of convergence is > 1 , we also have $G'(1) = E(X)$, $G''(1) = E(X^2) - E(X)$, $G'''(1) = E(X^3) - 3E(X^2) + 2E(X)$, \dots

X	$E(X)$	$\text{Var}(X)$	$G_X(z)$ or $\varphi_X(u)$
$\text{Bin}(n, p)$	np	$np(1-p)$	$(1-p+pz)^n$
$\text{Hyp}(N, K, n)$	Kn/N		
$\text{Poisson}(\mu)$	μ	μ	$\exp(\mu(z-1))$
$\text{Geom}(p)$	$(1-p)/p$	$(1-p)/p^2$	$p/(1-(1-p)z)$
$\text{Zeta}(\alpha)$	$\zeta(\alpha-1)/\zeta(\alpha)$ if $\alpha > 2$ ∞ otherwise		
$\text{Uniform}(a, b)$	$(a+b)/2$	$(b-a)^2/12$	
$\text{Gamma}(\alpha, \beta)$	α/β	α/β^2	
$\text{Normal}(\mu, \sigma^2)$	μ	σ^2	$\exp(i\mu u - \sigma^2 u^2/2)$

9. THE BOREL-CANTELLI LEMMA

Theorem 9.1. Let $(A_n)_{n < \omega}$ be a sequence of measurable sets in (X, \mathcal{X}, μ) .

- (1) If $\sum_{n < \omega} \mu(A_n) < \infty$, then $\mu(\limsup_n A_n) = 0$.
- (2) If $\mu = P$ is a probability, $\sum_{n < \omega} P(A_n) = \infty$, and the A_n 's are independent, then $P(\limsup_n A_n) = 1$.

10. STOCHASTIC PROCESSES

A family (even more than countable) of sub- σ -algebras $\{\mathcal{E}_i\}_{i \in I}$ of \mathcal{F} is P -independent if for every finite subset J of I and for every choice of $A_j \in \mathcal{E}_j$ we have

$$P\left(\bigcap_{j \in J} A_j\right) = \prod_{j \in J} P(A_j).$$

A family $\{X_i\}_{i \in I}$ of random variables is P -independent if the family $\{X_i^{-1}\mathcal{B}\}_{i \in I}$ is P -independent.

Definition 10.1. A stochastic process is a sequence of r.v.s $X_0, X_1, \dots : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B})$ (\mathbb{R} might be replaced by \mathbb{C} or \mathbb{R}^d). It is:

- independent if the X_n 's are independent;
- identically distributed if $(X_n)_*P = (X_m)_*P$ for every n and m ;
- stationary if the push-forward probability \bar{X}_*P on $\mathbb{R}^{\mathbb{Z}_{\geq 0}}$ is shift-invariant, that is, for every n and every choice of events A_0, \dots, A_{n-1} , we have

$$P(X_0 \in A_0 \cap \dots \cap X_{n-1} \in A_{n-1}) = P(X_1 \in A_0 \cap \dots \cap X_n \in A_{n-1}).$$

I.i.d. implies stationary, which implies i.d.. Let $\Omega = \{0, 1\}$ with uniform P . Let $X_n = \text{id}$ if $n \equiv 0 \pmod{3}$ and $X_n = 1 - \text{id}$ otherwise. Then $\bar{X}_*P([011]) = 1/2$ and $\bar{X}_*P([*011]) = 0$; thus \bar{X} is i.d. and non-stationary.

Lemma 10.2. Let g, g_0, g_1, \dots be measurable function.

- (1) If X_0, X_1, \dots is independent, then $g_0 \circ X_0, g_1 \circ X_1, \dots$ is independent.
- (2) If X_0, X_1, \dots is identically distributed, then $g \circ X_0, g \circ X_1, \dots$ is identically distributed.

Definition 10.3. For every n , let

$$\mathcal{C}_n = (\sigma\text{-algebra generated by } \bigcup_{m \geq n} X_m^{-1}\mathcal{B}) = \bigvee_{m \geq n} X_m^{-1}\mathcal{B}.$$

Then $\mathcal{C}_\infty = \bigcap_n \mathcal{C}_n$ is the *tail σ -algebra* of the process.

Lemma 10.4. Let $\{\mathcal{F}_i\}_{i \in I}$ be a family of independent σ -algebras. Let $\{I_j\}_{j \in J}$ be a partition of I and, for every $j \in J$, let $\mathcal{M}_j = \bigvee_{i \in I_j} \mathcal{F}_i$. Then the family $\{\mathcal{M}_j\}_{j \in J}$ is independent.

Proof. [Bil95, Theorem 4.2 p. 50] □

Theorem 10.5 (The Kolmogorov 0-1 Law). *The tail σ -algebra of an independent process is trivial.*

Example 10.6. Fix $p \in [0, 1]$, and let N_p be the set of all $x \in [0, 1]$ such that 5 appears in the decimal expansion of x with frequency p . Then either $\lambda(N_p) = 0$ or $\lambda(N_p) = 1$.

11. PRODUCT MEASURES

Definition 11.1. Let $(X, \mathcal{X}), (Y, \mathcal{Y})$ be measurable spaces. By Lemma 1.4, the class

$$\mathcal{S} = \{\pi_1^{-1}A \cap \pi_2^{-1}B : A \in \mathcal{X}, B \in \mathcal{Y}\} = \{A \times B : A \in \mathcal{X}, B \in \mathcal{Y}\}$$

is a semialgebra. We define $\mathcal{X} \times \mathcal{Y} = \mathcal{F}(\mathcal{S})$.

Lemma 11.2. Let X, Y be topological spaces with a countable basis, \mathcal{X}, \mathcal{Y} their Borel σ -algebras, \mathcal{B} the Borel σ -algebra of $X \times Y$. Then $\mathcal{B} = \mathcal{X} \times \mathcal{Y}$.

If $X, Y : \Omega \rightarrow \mathbb{R}$ are r.v.s, then $(X, Y) : \Omega \rightarrow \mathbb{R}^2$ is a r.v.. Since r.v.s are closed under postcomposition with continuous functions, the set of all r.v.s from Ω to \mathbb{R} is an \mathbb{R} -algebra.

Lemma 11.3. Let $f : X \times Y \rightarrow \mathbb{C}$ be $\mathcal{X} \times \mathcal{Y}$ -measurable. Then every $f(a, -)$ is \mathcal{Y} -measurable and every $f(-, b)$ is \mathcal{X} -measurable.

Let $C \in \mathcal{X} \times \mathcal{Y}$, and define

$$\begin{aligned} \varphi_C(x) &= \int_Y \mathbb{1}_C(x, -) d\nu \in [0, \infty], \\ \psi_C(y) &= \int_X \mathbb{1}_C(-, y) d\mu \in [0, \infty]. \end{aligned}$$

Theorem 11.4. φ_C is \mathcal{X} -measurable, ψ_C is \mathcal{Y} -measurable, and

$$\int_X \varphi_C d\mu = \int_Y \psi_C d\nu.$$

The function ρ that associates that number to C is a σ -finite measure, satisfies $\rho(A \times B) = \mu(A)\nu(B)$, and is the only measure on $(X \times Y, \mathcal{X} \times \mathcal{Y})$ that satisfies such an identity. We denote it by $\rho = \mu \times \nu$, and call it the *product measure* of μ and ν .

Theorem 11.5 (The Tonelli-Fubini Theorem). Let $(X, \mathcal{X}, \mu), (Y, \mathcal{Y}, \nu), (X \times Y, \mathcal{X} \times \mathcal{Y}, \mu \times \nu)$ be as above, and assume that $f : X \times Y \rightarrow \mathbb{C}$ is $\mathcal{X} \times \mathcal{Y}$ -measurable.

(1) If $f \geq 0$, then

$$\begin{aligned}\varphi_f(x) &= \int_Y f(x, -) d\nu \quad \text{is } \mathcal{X}\text{-measurable,} \\ \psi_f(y) &= \int_X f(-, y) d\mu \quad \text{is } \mathcal{Y}\text{-measurable,}\end{aligned}$$

and the identity

$$\int_{X \times Y} f d\mu \times \nu = \int_X \varphi_f d\mu = \int_Y \psi_f d\nu \quad (11.1)$$

holds.

(2) If any of the integrals in (11.1), with $|f|$ for f , is finite, then:

- (a) $f \in L_1(\mu \times \nu)$;
- (b) $f(x, -) \in L_1(\nu)$ for μ -every x , and $\varphi_f \in L_1(\mu)$;
- (c) $f(-, y) \in L_1(\mu)$ for ν -every y , and $\psi_f \in L_1(\nu)$;
- (d) the identity in (11.1) holds.

12. THE MULTIPLICATION THEOREM

Theorem 12.1. Let X, Y be r.v.'s and define $Z(\omega) = (X(\omega), Y(\omega))$. Then Z is a r.v., and X, Y are independent iff $Z_*P = X_*P \times Y_*P$.

Corollary 12.2. Let $X, Y \in L_1(P)$ be independent. Then $XY \in L_1(P)$ and $E(XY) = E(X)E(Y)$.

Corollary 12.3. Let X, Y be $\mathbb{Z}_{\geq 0}$ -valued and independent, and assume that G_X, G_Y have radius of convergence $\geq r > 1$. Then $G_{X+Y}(z) = G_X(z)G_Y(z)$ for $|z| < r$.

Taking into account the fact that a generating function determines the variable, this implies that the sum of two independent Poisson variables, of parameters μ and ν , is Poisson of parameter $\mu + \nu$.

Proposition 12.4. Let $X : (\Omega, P) \rightarrow \mathbb{R}$ and $X' : (\Omega', P') \rightarrow \mathbb{R}$ be r.v.s. Let $Y = X \circ \pi_1, Y' = X' \circ \pi_2 : (\Omega \times \Omega', \mathcal{F} \times \mathcal{F}') \rightarrow \mathbb{R}$. Then X and Y have the same law, X' and Y' have the same law, and Y, Y' are independent.

13. COVARIANCE

Let $X, Y \in L_2(P)$, $\mu = E(X), \nu = E(Y)$. The covariance of X and Y is

$$\text{Cov}(X, Y) = E(\overline{(X - \mu)(Y - \nu)}) \in \mathbb{C}.$$

Lemma 13.1. (1) $\text{Var}(X) = \text{Cov}(X, X)$;

(2) Cov is hermitian sesquilinear;

(3) $\text{Cov}(X, Y) = E(\overline{XY}) - \bar{\mu}\nu$;

(4) if X, Y are independent, their covariance is 0;

(5) the variance of a finite sum of independent variables is the sum of the variances.

The correlation coefficient of X, Y is

$$\rho = \frac{\text{Cov}(X, Y)}{(\text{Var}(X) \text{Var}(Y))^{1/2}}.$$

Lemma 13.2. ρ belongs to the closed unit disc (it is in $[-1, 1]$ if X, Y are real-valued), and equals 0 if X, Y are independent.

Proof.

$$|\text{Cov}(X, Y)| = \left| E(\overline{(X - \mu)}(Y - \nu)) \right| \leq [E(|X - \mu|^2) E(|Y - \nu|^2)]^{1/2}$$

by Cauchy-Schwarz. □

Definition 13.3. Let $X = (X_1, \dots, X_d) : \Omega \rightarrow \mathbb{C}^d$ be a r.v., with all components in $L_2(P)$. Its covariance matrix is the hermitian-symmetrix matrix Σ^2 whose ij th entry is $\text{Cov}(X_i, X_j)$.

Lemma 13.4. Let $S : \mathbb{C}^d \rightarrow \mathbb{C}^m$ be a linear map. Then the covariance matrix of $S \circ X$ is $\overline{S} \Sigma^2 S^\top$. In particular, Σ^2 is positive semidefinite.

14. DENSITIES

Lemma 14.1. Let $m : \mathbb{R}^d \rightarrow \mathbb{R}_{\geq 0}$ be in $L_1(\lambda)$, with $\int m \, d\lambda = 1$. Then the function $\mu_m : \mathcal{B} \rightarrow [0, 1]$ defined by

$$\mu_m(A) = \int_{\mathbb{R}^d} \mathbb{1}_A m \, d\lambda$$

is a probability. The map $m \mapsto \mu_m$ is injective.

If μ is in the range of the above map, then we say that μ has a density (which is unique).

Lemma 14.2. Let $f : \mathbb{R}^d \rightarrow \mathbb{C}$ be measurable. Then $f \in L_1(\mu_m)$ iff $fm \in L_1(\lambda)$. If this happens, then

$$\int_{\mathbb{R}^d} f \, d\mu_m = \int_{\mathbb{R}^d} fm \, d\lambda.$$

Let $A : \mathbb{R} \rightarrow \mathbb{R}$ be a homeomorphism, μ a probability on \mathbb{R} with repartition function M . Then $A_*\mu$ has repartition function $M \circ A^{-1}$ if A is increasing, and $1 - M(A^{-1}x - 0)$ if A is decreasing.

Let $T : I \rightarrow \mathbb{R}$, with I an interval in \mathbb{R} and T piecewise C^1 and strictly monotone. Let μ be the probability on I induced by the continuous density $m : I \rightarrow \mathbb{R}_{\geq 0}$.

Theorem 14.3. Under the above hypotheses $T_*\mu$ has a density $\mathcal{L}m$, which is explicitly given by

$$\begin{aligned} (\mathcal{L}m)(y) &= \sum \left\{ \frac{m(x)}{|T'(x)|} : x \in T^{-1}\{y\} \right\} \\ &= \sum \{ |A'(y)| m(A(y)) : A \text{ is an inverse branch of } T \text{ and } y \in \text{dom}(A) \}. \end{aligned}$$

The map \mathcal{L} is the *Ruelle-Perron-Frobenius operator*, or *transfer operator*; more generally, Theorem 14.3 holds for $m \in L_1(\lambda)$.

Theorem 14.4. Let $T : O \rightarrow \mathbb{R}^d$, with O open in \mathbb{R}^d and T injective C^1 with never 0 jacobian determinant j_T . Assume that the probability μ on O is induced by the density $m \in L_1(\lambda)$, and let $A = T^{-1}$. Then $T_*\mu$ is induced by a density, which is explicitly given by

$$\mathcal{L}m = \frac{m}{|j_T|} \circ T^{-1} = |j_A|(m \circ A)$$

on $T[O]$, and 0 otherwise. There's an analogous statement for the piecewise case.

15. MARGINALS

Definition 15.1. If P is a probability on $(\Omega \times \Omega', \mathcal{F} \times \mathcal{F}')$, then $\pi_{1*}P, \pi_{2*}P$ are the *marginals* of P . Conversely, any P that projects to a given pair P_1, P_2 is a *joining* of P_1 and P_2 .

Theorem 15.2. Let μ be a probability on \mathbb{R}^2 with marginals μ_1, μ_2 , and assume that μ has a density m .

(1) μ_1 and μ_2 have densities, explicitly given by

$$m_1(x) = \int_{\mathbb{R}} m(x, -) d\lambda,$$

$$m_2(y) = \int_{\mathbb{R}} m(-, y) d\lambda.$$

(2) π_1 and π_2 are μ -independent iff $m(x, y) = m_1(x)m_2(y)$ in $L_1(\lambda^2)$.

(3) The set $A = \{a \in \mathbb{R} : m_1(a) \neq 0, +\infty\}$ has full μ_1 -measure, and parametrizes a family of densities on \mathbb{R} , namely

$$m(y|a) = \frac{m(a, y)}{m_1(a)}.$$

(4) For every $a \in A$, let $(\mu|a)$ be the pushforward via ι_a of the probability on \mathbb{R} of density $m(-|a)$; in other words,

$$(\mu|a)(B) = \int_{\mathbb{R}} \mathbb{1}_B(a, y)m(y|a) d\lambda(y).$$

We then have

$$\mu = \int_{\mathbb{R}} (\mu|x) d\mu_1(x),$$

which is the continuous version of the first Bayes identity.

16. THE GAUSSIAN IN \mathbb{R}^d

The standard normal distribution is the distribution of a random variable $Z = (Z_1, \dots, Z_d) : \Omega \rightarrow \mathbb{R}^d$ with Z_1, \dots, Z_d independent standard normals. In other words, Z_*P is induced by the density

$$m(x_1, \dots, x_d) = \frac{1}{\sqrt{(2\pi)^d}} \exp\left(-\frac{1}{2}(x_1^2 + \dots + x_d^2)\right) = \frac{1}{\sqrt{(2\pi)^d}} \exp\left(-\frac{1}{2}\langle x, x \rangle\right).$$

Applying $T(x) = Sx + \mu : \mathbb{R}^d \rightarrow \mathbb{R}^d$ we get the general case $T \circ Z \in \text{Normal}(\mu, \Sigma^2)$. It has density

$$m(x) = \frac{1}{\sqrt{(2\pi)^d \det(\Sigma^2)}} \exp\left(-\frac{1}{2}(x - \mu)^T \Sigma^{-2}(x - \mu)\right),$$

where $\Sigma^2 = SS^T$ is the covariance matrix, which is positive definite.

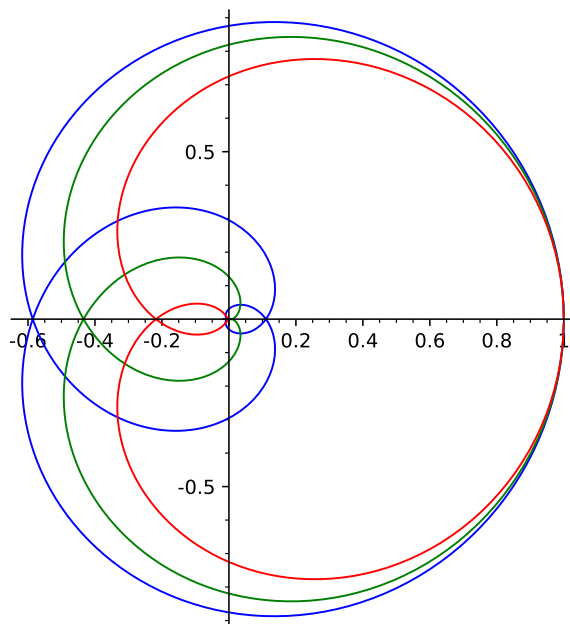


FIGURE 6

17. CHARACTERISTIC FUNCTIONS

Definition 17.1. Let μ be a probability on \mathbb{R}^d . Its *Fourier transform* is the function $\hat{\mu} : \mathbb{R}^d \rightarrow \mathbb{C}$ given by

$$\hat{\mu}(s) = \int_{\mathbb{R}^d} \exp(i\langle s, x \rangle) d\mu(x).$$

If $\mu = X_*P$, then $\hat{\mu}$ is the *characteristic function* of X , written

$$\varphi_X(s) = E(\exp(i\langle s, X \rangle)).$$

If X is $\mathbb{Z}_{\geq 0}$ -valued, then $\varphi_X(s) = G_X(\exp(is))$, which is 2π -periodic.

Theorem 17.2. $\hat{\mu}$ is uniformly continuous, bounded by $\mathbb{1}$, and $\hat{\mu}(0) = 1$.

Theorem 17.3. (1) $\varphi_{aX} = \varphi_X(a-)$;
 (2) $\varphi_{-X} = \overline{\varphi_X}$;
 (3) if $X = -X$ in law, then φ_X is \mathbb{R} -valued;
 (4) if X and Y are independent, then $\varphi_{X+Y} = \varphi_X\varphi_Y$.

It is not true that $\varphi_{X+Y} = \varphi_X\varphi_Y$ implies that X and Y are independent [JP03, p. 113].

Example 17.4. If X is *Uniform* $(-a, a)$, then $\varphi_X(s) = \sin(as)/(as)$.

Theorem 17.5 (The Theorem of Moments). Let $X : \Omega \rightarrow \mathbb{R}$ be in $L_m(P)$, for some $m \geq 1$. Then $\varphi_X \in C^m(\mathbb{R})$ and

$$\varphi_X^{(m)}(s) = E((iX)^m \exp(isX)).$$

In particular, $E(X^m) = (-i)^m \varphi_X^{(m)}(0)$.

It has a d -dimensional version.

Theorem 17.6. *Let $X : \Omega \rightarrow \mathbb{R}^d$, and let $m \geq 1$ be such that, for every $1 \leq r \leq m$ and every j_1, \dots, j_r , the product $X_{j_1} \cdots X_{j_r}$ is in $L_1(P)$. Then $\varphi_X \in C^m(\mathbb{R}^d)$ and*

$$(\partial_{s_{j_1}} \cdots \partial_{s_{j_m}} \varphi_X)(s) = E((iX_{j_1}) \cdots (iX_{j_m}) \exp(i\langle s, X \rangle)).$$

Theorem 17.7. *The map $\mu \mapsto \hat{\mu}$ is injective.*

Proof. [JP03, Theorem 14.1] □

Corollary 17.8. (1) $\varphi_X = \varphi_Y$ iff X and Y have the same law.

(2) $X = -X$ in law iff φ_X is \mathbb{R} -valued;

Lemma 17.9. *Let X be a standard normal. Then $\varphi_X(s) = \exp(-s^2/2)$.*

18. CONVOLUTIONS

Definition 18.1. Let $\alpha : M \times X \rightarrow X$ be a left action of a monoid M on a set X . Let μ and ν be probabilities on M and X , respectively. Then $\mu * \nu = \alpha_*(\mu \times \nu)$ is the *convolution product* of μ and ν , which yields a left action of $\mathcal{P}(M)$ on $\mathcal{P}(X)$.

The main example is that of a group G acting on itself by left translations; we will deal with the case $G = (\mathbb{R}, +)$ only.

Theorem 18.2. (1) $\widehat{\mu * \nu} = \hat{\mu}\hat{\nu}$.

(2) If X, Y are independent then $(X_*P) * (Y_*P) = (X + Y)_*P$, and therefore $\varphi_{X+Y} = \varphi_X\varphi_Y$.

Theorem 18.3. *Let μ, ν be probabilities on $(\mathbb{R}, \mathcal{B})$, and assume that μ has density m . Then $\mu * \nu$ has density r , which is given by*

$$r(z) = \int_{\mathbb{R}} m(z - y) d\nu(y).$$

If ν has density n as well, then

$$r(z) = \int_{\mathbb{R}} m(z - y)n(y) d\lambda(y),$$

which defines the convolution $r = m * n$ of the densities m and n .

19. CONVERGENCE OF R.V.S

Definition 19.1. Let $f, f_0, f_1, f_2, \dots : (X, \mathcal{X}, \mu) \rightarrow \mathbb{C}$ be measurable. If, for every $\varepsilon > 0$, $\mu(|f_n - f| > \varepsilon)$ converges to 0, then we say that f_n converges to f *in measure*.

Lemma 19.2. *Let $h : \mathbb{R}_{\geq 0} \rightarrow [0, M]$ be such that $h(0) = 0$, h is continuous nondecreasing, and strictly increasing in some right neighborhood of 0. Then $X_n \rightarrow X$ in probability iff $E(h \circ |X_n - X|) \rightarrow 0$.*

Theorem 19.3. *Fix $1 \leq p < \infty$.*

(1) *If X_n converges to X either a.e. or in L_p , then it converges in probability.*

(2) *If X_n converges to X in probability, then there exists a subsequence that converges a.e.*

(3) If X_n converges to X in probability, and there exists $0 \leq Y \in L_p$ that dominates every X_n , then X_n converges to X in L_p .

Given a stochastic process X_0, X_1, \dots , we set

$$S_n = \sum_{k < n} X_k, \quad A_n = \frac{S_n}{n}.$$

Remark 19.4. The process S_0, S_1, \dots is the *random walk* induced by X on the group \mathbb{R} . More generally, if X_0, X_1, \dots take values in a not necessarily commutative semigroup, then $L_n = X_{n-1} \cdots X_0$ is the induced *left random walk* and $R_n = X_0 \cdots X_{n-1}$ the right one.

Theorem 19.5 (The Weak Law of Large Numbers). *Assume the process is i.i.d., with all variables in $L_2(P)$; let $\mu = E(X_0)$. Then $A_n \rightarrow \mu \mathbb{1}$ in $L_2(P)$ and in probability.*

20. WEAK CONVERGENCE

Let μ, μ_0, μ_1, \dots be probabilities on \mathbb{R}^d . Let $C_b(\mathbb{R}^d)$ be the set of all continuous bounded functions from \mathbb{R}^d to \mathbb{R} (or \mathbb{C}). If, for every $f \in C_b(\mathbb{R}^d)$,

$$\int f \, d\mu_n \rightarrow \int f \, d\mu,$$

then we say that μ_n converges to μ *weakly*. If $(X_n)_*P$ converges to X_*P weakly, then we say that X_n converges to X *weakly* (or *in distribution*, or *in law*); this amounts to $E(f \circ X_n) \rightarrow E(f \circ X)$, for every $f \in C_b(\mathbb{R}^d)$.

Remark 20.1. One can replace $C_b(\mathbb{R}^d)$ with its subset of bounded Lipschitz functions [JP03, Theorem 18.7], or with any subset whose \mathbb{R} -span is uniformly dense.

Example 20.2. A sequence r_0, r_1, \dots in $[0, 1]$ is *uniformly distributed* w.r.t. the Lebesgue measure λ if the sequence of Cesàro averages $n^{-1} \sum_{k=0}^{n-1} \delta_{r_k}$ converges weakly to λ . The basic example is $r_n = \alpha_1 n + \alpha_0 \pmod{1}$, with α_1, α_0 real numbers and α_1 irrational. This is a first instance of the Weyl equidistribution theorem, and can be proved by using as test functions the family $\chi_k = \exp(2\pi i k -)$, for $k \in \mathbb{Z}$, whose \mathbb{C} -span is uniformly dense in $C(\mathbb{R}/\mathbb{Z}, \mathbb{C})$.

Theorem 20.3. (1) *Convergence in probability implies weak convergence.*
 (2) *Weak convergence to a constant implies convergence in probability.*

Theorem 20.4. *Let μ, μ_0, μ_1, \dots be probabilities on \mathbb{R} , and let M, M_0, M_1, \dots be their repartition functions. Then $\mu_n \rightarrow \mu$ weakly iff $M_n \rightarrow M$ at every point at which M is continuous.*

Theorem 20.5. *Let μ, μ_0, μ_1, \dots be probabilities on a finite or countable space, and let m, m_0, m_1, \dots be their discrete-density functions. Then $\mu_n \rightarrow \mu$ weakly iff $m_n \rightarrow m$ pointwise.*

Example 20.6. Let p_n go to 0 as n goes to infinity, in such a way that np_n converges to some constant $\mu \in \mathbb{R}_{>0}$. Then the sequence $\text{Bin}(n, p_n)$ weakly converges to $\text{Poisson}(\mu)$. Thus, a binomial with large n and small p can be effectively (i.e., the discrete-density function is more manageable) approximated by a Poisson of parameter np .

Definition 20.7. A family $\{\mu_i\}_{i \in I}$ of probabilities on \mathbb{R}^d is *tight* if for every $\varepsilon > 0$ there exists a compact $K \subset \mathbb{R}^d$ such that, for every i , $\mu_i(K^c) < \varepsilon$.

Theorem 20.8 (The Helly Selection Theorem). *Every sequence extracted from a tight family contains a weakly converging subsequence.*

Theorem 20.9 (Slutski's Theorem). *Let $X_0, X_1, \dots, Y_0, Y_1, \dots, Z : \Omega \rightarrow \mathbb{R}^d$. Let $\|\cdot\|$ be any norm on \mathbb{R}^d . Assume that $X_n \rightarrow Z$ weakly and $\|X_n - Y_n\| \rightarrow 0$ in probability. Then $Y_n \rightarrow Z$ weakly.*

Theorem 20.10 (The Lévy Continuity Theorem). *Let $(\mu_n)_{n < \omega}$ be a sequence of probabilities on \mathbb{R}^d . Then:*

- (1) *if $\mu_n \rightarrow \mu$ weakly, then $\hat{\mu}_n \rightarrow \hat{\mu}$ everywhere (actually, uniformly on compacta);*
- (2) *if $\hat{\mu}_n \rightarrow f$ everywhere and f is continuous at 0, then $f = \hat{\mu}$ for some μ and $\mu_n \rightarrow \mu$ weakly.*

21. THE STRONG LAW OF LARGE NUMBERS

Theorem 21.1. *Assume X_0, X_1, \dots is i.i.d., with all variables in $L_2(P)$; let $\mu = E(X_0)$. Then $A_n \rightarrow \mu \mathbb{1}$ a.e.*

This yields, e.g., Monte Carlo integration and Borel's normal number theorem.

22. THE CENTRAL LIMIT THEOREM

Theorem 22.1. *Let $X_0, X_1, \dots : \Omega \rightarrow \mathbb{R}$ be i.i.d., with variables in $L_2(P)$ having mean μ and variance $\sigma^2 > 0$. Let*

$$Y_n = \frac{S_n - n\mu}{\sigma\sqrt{n}} = \frac{A_n - \mu}{\sigma/\sqrt{n}}.$$

Then Y_n converges weakly to a standard normal.

Example 22.2. We want to approximate $\pi/4$, namely the area of a disk of radius $1/2$ inscribed in the unit square, by firing bullets randomly. We need to compute the minimum n such that with probability 99% our shooting sequence will approximate the area up to $d \geq 1$ correct decimal places. We have $\mu = \pi/4$, $\sigma^2 = \mu - \mu^2$, and thus require

$$P(|A_n - \mu| < 10^{-d}) = P\left(\frac{|A_n - \mu|}{\sigma/\sqrt{n}} < \frac{10^{-d}}{\sigma/\sqrt{n}}\right) \approx P(|Z| < (\sqrt{n}/\sigma)10^{-d}) \geq 0.99.$$

The equality \approx above is not mathematically rigorous, but works well in practice. By looking at tables, or using a calculator, this happens for $(\sqrt{n}/\sigma)10^{-d} > \sqrt{2} \cdot 1.183$, i.e., $n > 2\sigma^2 1.83^2 100^d$; thus 11289 bullets for 2 digits.

Example 22.3. Let the functions $X_2, X_3, X_5, \dots : \mathbb{N} \rightarrow \mathbb{R}_{\geq 0}$ be defined by $X_p(n) = 1$ if $p \mid n$, and 0 otherwise. It is an astonishing fact that the X_p s behave as if they were an independent process. This provides heuristics for facts that can often be proved by non-probabilistic means; here's an easy example and a difficult one.

- (1) Fix $d \geq 2$, let l be a large number and let (a_1, \dots, a_d) vary in $\{1, \dots, l\}^d$. For $p \leq l$ we have $P(X_p(a_1) = \dots = X_p(a_d) = 1) = p^{-d}$. The above heuristics gives then

$$P(a_1, \dots, a_d \text{ are relatively prime}) = \prod_{p \leq l} (1 - p^{-d}),$$

which tends to $\zeta(d)^{-1}$ for l tending to infinity.

- (2) Let $\omega(n) = \sum_p X_p(n)$ be the number of prime divisors of n , and let m be the density function of a standard normal. Then the Erdős-Kac theorem says that, for every a ,

$$\lim_{l \rightarrow \infty} \frac{1}{l} \#\left\{n \leq l : \frac{\omega(n) - \log \log l}{\sqrt{\log \log l}} \leq a\right\} = \int_{-\infty}^a m(x) dx.$$

Thus if we pick a number at random in $\{1, \dots, 10^{10000}\}$, the number of its prime factors will be gaussian distributed with average and variance 10.044. Therefore, since 90% of the mass of a standard normal lies between -1.644 and 1.644 , with probability 90% the number of distinct prime factors of our number will be between 4.83 and 15.25.

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