



Asymptotic Behaviour of a q -Binomial Type Distribution Based on q -Krawtchouk Orthogonal Polynomials

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ABSTRACT

In this manuscript, we introduce for $0 < q < 1$ a new deformed q -binomial distribution by virtue of the normalized q -Krawtchouk orthogonal polynomials. We call this q -discrete probability distribution q -binomial distribution of type I, say $\mathcal{B}(n, p, q)$ with $p > 0$. Interpretation of this q -discrete probability distribution is provided by considering, for suitable values of q , n -dimensional vector spaces over the finite field of $1/q$ elements. Furthermore, we study its asymptotic behaviour for $n \rightarrow \infty$ by showing that is compatible with a standardized Stieltjes-Wigert distribution.

1. INTRODUCTION

Let v_n be a discrete probability measure in $S = \{x_i, i = 1, 2, \dots, n; n = 0, 1, \dots\}$ with finite moments of all orders. Then it is well known that there exist a sequence of normalized discrete orthogonal polynomials $\{P_m(x)\}$ with respect to the measure v_n satisfying the orthogonality relation

$$\sum_{x \in S} v_n(x) P_m(x) P_\nu(x) = \lambda_{n\nu} \delta_{m,\nu}, \quad (1)$$

where $\delta_{m,\nu}$ the Kronecker delta and $\lambda_{n\nu}$ a non-negative sequence of n and ν and the three term recurrence relation

$$x P_m(x) = P_{m+1}(x) + a_m P_m(x) + b_m P_{m-1}(x) \quad (m \geq 1), \quad (2)$$

where $a_m \in \mathbb{R}$ and $b_m > 0$ and with initial conditions $P_0(x) = 1$ and $P_1(x) = x - a_0$. Conversely, Favard's theorem ensures the existence of a discrete probability measure v_n on S for which the sequence of polynomials

determined by the recurrence relation (2) are orthogonal. The mean value and the variance of the discrete random variable X in the discrete spectrum S with probability function $v_n(x)$ are given respectively by $\mu = a_0$ and $\sigma^2 = b_1$. If $a_m = 0$ then all moments of odd order are zero.(see Saitoh and Yoshida (2000), Christiansen (2004)).

Saitoh and Yoshida(2000) introduced a q -deformed binomial distribution for $0 < q < 1$, by virtue of a q -deformed sequence of Krawtchouk orthogonal polynomials and studied its asymptotic behaviour by showing that is compatible with a q -deformed Gaussian distribution in a quantum probability space.

Various q -analogues for $0 < q < 1$, of the classical binomial distribution have also been studied by many authors. Among them we refer to Kemp(1992, 2002), Sicong(1994) and Charalambides(2005).

In this manuscript, we introduce for $0 < q < 1$ a new deformed q -binomial distribution by virtue of the normalized q -Krawtchouk orthogonal polynomials. Interpretation of this q -discrete probability distribution is provided by considering, for suitable values of q , n -dimensional vector spaces over the finite field of $1/q$ elements. Furthermore, we study its asymptotic behaviour for $n \rightarrow \infty$ by showing that is compatible with a standardized Stieltjes-Wigert distribution.

2.INTRODUCTORY DEFINITIONS AND NOTATIONS

For the needs of this manuscript we recall some usual definitions and notation used in q -analysis (see Koekoek and Swarttouw (1998)).

Let $0 < q < 1$, x a real number and k a positive integer. The q -shifted factorial is defined by

$$(a; q)_n := \prod_{j=1}^n (1 - aq^{j-1})$$

and the general q -shifted factorial is given by

$$(a; q)_\infty := \prod_{j=1}^{\infty} (1 - aq^{j-1}), \quad (a; q)_0 = 1.$$

Also,

$$(a, b; q)_n := (a; q)_n (b; q)_n.$$

The q -number is defined by

$$[x]_q = \frac{1 - q^x}{1 - q}$$

and the q -binomial coefficient by

$$\binom{n}{k}_q := \frac{(q; q)_n}{(q; q)_k (q; q)_{n-k}}$$

or

$$\binom{n}{k}_{1/q} = q^{-k(n-k)} \binom{n}{k}_q.$$

Also

$$(-t; q)_n = \sum_{k=0}^n \binom{n}{k}_q q^{\binom{k}{2}} t^k.$$

3. q -BINOMIAL DISTRIBUTION OF TYPE I-ASYMPTOTIC BEHAVIOUR

Let $0 < q < 1$ be a rational number such that $1/q$ be an integer larger than 1. Also, let V_n be an n -dimensional vector space over the finite field $GF(1/q)$ of $1/q$ elements. Then V_n contains a total of $1/q^n$ vectors and the number of x -dimensional subspaces of V_n , $x = 0, 1, \dots, n$ is given by $\binom{n}{x}_{1/q}$ (see Exton(1983)).

Supposing that each x -dimensional subspace assigns a weight, say $g_q(x; V_n)$, with

$$g_q(x; V_n) = q^{\binom{x+1}{2}} p^{-x}, \quad p > 0, \quad x = 0, 1, \dots, n$$

we have that the probability of appearance of a x -dimensional subspace of V_n is given by

$$f_X(x) = \frac{\binom{n}{x}_{1/q} q^{\binom{x+1}{2}} p^{-x}}{\sum_{x=0}^n \binom{n}{x}_{1/q} q^{\binom{x+1}{2}} p^{-x}} \quad (3)$$

or

$$f_X(x) = \binom{n}{x}_q q^{\binom{x}{2}} (p^{-1} q^{-n})^x \prod_{i=1}^n (1 + p^{-1} q^{i-n-1})^{-1}, \quad x = 0, 1, \dots, n, \quad (4)$$

with $0 < q < 1$, $p > 0$.

Next, we define the following deformation of the q -analogue binomial distribution (4) considering the random variable $Y = q^{-X}$ with probability function

$$u_Y(y) = \frac{\binom{n}{-\frac{\ln y}{\ln q}}_q q^{\binom{-\frac{\ln y}{\ln q}}{2}} \alpha_n^{-\frac{\ln y}{\ln q}}}{\prod_{i=1}^n (1 + \alpha_n q^{i-1})}, \quad y = q^{-0}, q^{-1}, \dots, q^{-n}, \quad (5)$$

where

$$\alpha_n = p^{-1}q^{-n}, \quad p > 0 \tag{6}$$

such that

$$\alpha_n > 1, \quad \alpha_n q^n \rightarrow 0 \tag{7}$$

and

$$\alpha_n \rightarrow \infty \text{ as } n \rightarrow \infty. \tag{8}$$

Proposition 3.1. *The probability function $u_Y(y)$ is induced by the normalized q -Krawtchouk orthogonal polynomials, say $p_m(y; p, n, q)$, $0 < q < 1$ given by*

$$p_m(y; p, n, q) = \frac{(q^{-n}; q)_m}{(-pq^m; q)_m} K_m(y; p, n, q), \tag{9}$$

where $y = q^{-x}$, $x = 0, 1, \dots, n$ and $K_m(y; p, n, q)$ the q -Krawtchouk orthogonal polynomials with parameter $p = \alpha_n^{-1}q^{-n}$.

Moreover the random variable Y has the following r th order moments

$$\mu_r = \sum_{y \in S_1} y^r u_Y(y) = q^{-rn} \frac{\prod_{i=1}^r (1 + \alpha_n^{-1}q^i)}{\prod_{i=1}^r (1 + \alpha_n^{-1}q^{-n+i})}. \tag{10}$$

Proof. The normalized q -Krawtchouk orthogonal polynomials $p_m(q^{-x}; p, n, q)$, $0 < q < 1$ with parameter $p = \alpha_n^{-1}q^{-n}$ satisfies the orthogonality relation

$$\begin{aligned} & \sum_{x=0}^n \frac{(q^{-n}; q)_x}{(q; q)_x} (-1)^x \alpha_n^x q^{nx} p_m(q^{-x}; p, n, q) p_\nu(q^{-x}; p, n, q) \\ &= \gamma_{n\nu} (1 + \alpha_n^{-1}q^{-n}) (-\alpha_n^{-1}q^{-n+1}; q)_n \alpha_n^n q^{n^2} q^{-\binom{n+1}{2}} \delta_{m\nu} \end{aligned}$$

where $\delta_{m\nu}$ the Kronecker delta and $\gamma_{n\nu}$ a non-zero sequence of n and ν (see Koekoek and Swarttouw(1998)). Thus their weight function, say $u(x; n, q)$, can be written as

$$u(x; n, q) = \frac{\frac{(q^{-n}; q)_x}{(q; q)_x} (-1)^x \alpha_n^x q^{nx}}{(1 + \alpha_n^{-1}q^{-n}) (-\alpha_n^{-1}q^{-n+1}; q)_n \alpha_n^n q^{n^2} q^{-\binom{n+1}{2}}}. \tag{11}$$

Using the identities

$$\frac{(q^{-n}; q)_x}{(q; q)_x} = (-1)^x q^{-nx} \alpha_n^{-x} \binom{n}{x}_q q^{\binom{x}{2}}$$

and

$$(1 + \alpha_n^{-1} q^{-n})(-\alpha_n^{-1} q^{-n+1}; q)_n = \alpha_n^{-n} q^{-n^2} q^{\binom{n+1}{2}} \prod_{i=1}^n (1 + \alpha_n q^i)$$

we have from (11)

$$u(x; n, q) = u_Y(y), \quad y = q^{-x}.$$

Moreover, by the probability function (5) the r th moment of the random variable Y is

$$\begin{aligned} \mu_r &= \sum_{y \in S_1} y^r u_Y(y) \\ &= \frac{\sum_{y \in S_1} y^r \left(-\frac{n}{\ln q}\right)_q q^{\left(-\frac{\ln y}{2}\right)} \alpha_n^{-\frac{\ln y}{\ln q}}}{\prod_{i=1}^n (1 + \alpha_n q^{i-1})} \\ &= \frac{\sum_{x=0}^n q^{-xr} \binom{n}{x}_q q^{\binom{x}{2}} \alpha_n^x}{\prod_{i=1}^n (1 + \alpha_n q^{i-1})} \\ &= \frac{\prod_{i=1}^n (1 + \alpha_n q^{i-r-1})}{\prod_{i=1}^n (1 + \alpha_n q^{i-1})} \\ &= \frac{(1 + \alpha_n q^{-r})(1 + \alpha_n q^{-r+1}) \cdots (1 + \alpha_n^{-1} q^{-1}) \cdots (1 + \alpha_n q^{n-r-1})}{(1 + \alpha_n q^{-1}) \cdots (1 + \alpha_n q^{n-r+1})(1 + \alpha_n q^{n-r}) \cdots (1 + \alpha_n q^{n-1})} \end{aligned} \quad (12)$$

From the last expression the equation (10) is obtained.

Definition 3.1. For $0 < q < 1$, the q -discrete probability distribution with probability function $u_Y(y)$ defined in the spectrum $S_1 = \{q^{-k}, k = 0, 1, \dots, n\}$ and based on normalized q -Krawtchouk orthogonal polynomials is called q -Binomial of type I distribution and is denoted by $\mathcal{B}(n, p, q)$.

Remark 3.1. Using the moments of r th order (10) the random variable Y of the q -binomial of type I distribution has respectively the following mean value and variance

$$\mu_Y = q^{-n} \frac{1 + \alpha_n^{-1} q}{1 + \alpha_n^{-1} q^{-n+1}} \quad (13)$$

and

$$\sigma_Y^2 = q^{-2n+1} \alpha_n^{-1} \frac{(1 + \alpha_n^{-1} q)(q^{-n} - 1)(1 - q)}{(1 + \alpha_n^{-1} q^{-n+1})^2 (1 + \alpha_n^{-1} q^{-n+2})} \quad (14)$$

Equivalently, since the q -binomial of type I distribution is induced from the normalized q -Krawtchouk orthogonal polynomials $p_m(y; p, n, q)$, $y = q^{-x}$, $p = \alpha_n^{-1}q^{-n}$, the mean value and variance of the random variable Y can be derived from the recurrence relation of these polynomials. Analytically, the recurrence relation of the orthogonal polynomials $p_m(y; p, n, q)$ with $p = \alpha_n^{-1}q^{-n}$ is

$$\begin{aligned}
 yp_m(y; p, n, q) &= p_{m+1}(y; p, n, q) + a_m^{(p,n,q)}p_m(y; p, n, q) \\
 &+ b_m^{(p,n,q)}p_{m-1}(y; p, n, q) \quad (m \geq 1)
 \end{aligned}
 \tag{15}$$

with initial conditions

$$p_0(y; p, n, q) = 1, \quad p_1(y; p, n, q) = y - a_0,$$

where

$$a_m^{(p,n,q)} = 1 - (A_m + C_m), \quad b_m^{(p,n,q)} = A_{m-1}C_m \tag{16}$$

and

$$A_m = \frac{(1 - q^{m-n})(1 + \alpha_n^{-1}q^{m-n})}{(1 + \alpha_n^{-1}q^{2m-n})(1 + \alpha_n^{-1}q^{2m-n+1})}$$

(17)

$$C_m = -\alpha_n^{-1}q^{2m-2n-1} \frac{(1 + \alpha_n^{-1}q^{m-n})(1 - q^m)}{(1 + \alpha_n^{-1}q^{2m-n-1})(1 + \alpha_n^{-1}q^{2m-n})}$$

(see Koekoek and Swarttouw (1998)).

So,

$$\mu_Y = 1 - A_0 - C_0, \quad \sigma_Y^2 = A_0C_1$$

and on using (17) we recapture (13) and (14).

Theorem 3.1. *The limit distribution for $n \rightarrow \infty$ of the standardized q -binomial of type I distribution $\mathcal{B}(n, p, q)$ is the standardized Stieltjes-Wigert distribution with probability density function*

$$\begin{aligned}
 u_q^{SW}(y) &= \frac{q^{-11/8}(q^{-3/2}(1-q)^{1/2}y + q^{-1})^{-1/2}}{\sqrt{2\pi \log q^{-1}}} e^{\frac{\log(q^{-3/2}(1-q)^{1/2}y + q^{-1})^2}{2 \log q}}, \\
 & \qquad \qquad \qquad y > -q^{1/2}(1-q)^{-1/2}. \tag{18}
 \end{aligned}$$

Proof. Shifting so as to have zero mean in the q -Binomial of type I distribution with mean value μ_Y , its orthogonal polynomials are $p_m(y + \mu_Y; p, n, q)$,

$p = \alpha_n^{-1}q^{-n}$. The coefficients of the recurrence relation of $p_m(y + \mu_Y; p, n, q)$ are represented as $a_m^{(n,p,q)} - \mu_Y = A_0 + C_0 - (A_m + C_m)$ and $b_m^{(n,p,q)} = A_{m-1}C_m$. In addition, to standardized so as to be of variance 1, it can be realized by replacing the recurrence relation coefficients respectively by

$$\frac{a_m^{(n,p,q)} - \mu_Y}{\sigma_Y} = \frac{A_0 + C_0 - (A_m + C_m)}{\sqrt{A_0C_1}} \quad (19)$$

and

$$\frac{b_m^{(n,p,q)}}{\sigma_Y^2} = \frac{A_{m-1}C_m}{A_0C_1}, \quad (20)$$

where A_m and C_m are defined in (17).

Thus the orthogonal polynomials, say $Q_m(y; p, q, n)$, $p = \alpha_n^{-1}q^{-n}$, of the standardized q -binomial of type I distribution are determined by the recurrence relation

$$\begin{aligned} yQ_m(y; p, q, n) &= Q_{m+1}(y; p, q, n) + \frac{A_0 + C_0 - (A_m + C_m)}{\sqrt{A_0C_1}}Q_m(y; p, q, n) \\ &+ \frac{A_{m-1}C_m}{A_0C_1}Q_{m-1}(y; p, q, n), \quad (m \geq 1) \end{aligned} \quad (21)$$

with initial conditions

$$Q_0(y; p, q, n)(y) = 1 \quad Q_1(y; p, q, n)(y) = y.$$

Using (13), (14) and (17) the coefficients (19) and (20) respectively become

$$\begin{aligned} \frac{a_m^{(n,p,q)} - \mu_Y}{\sigma_Y} &= \\ &1 - \frac{(1-q^{m-n})(1+\alpha_n^{-1}q^{m-n})}{(1+\alpha_n^{-1}q^{2m-n})(1+\alpha_n^{-1}q^{2m-n+1})} + \frac{\alpha_n^{-1}q^{2m-2n-1}(1+\alpha_n^{-1}q^{m-n})(1-q^m)}{(1+\alpha_n^{-1}q^{2m-n-1})(1+\alpha_n^{-1}q^{2m-n})} \\ &\frac{q^{-n+1/2}\alpha_n^{-1/2}(1+\alpha_n^{-1}q)^{1/2}(q^{-n}-1)^{1/2}(1-q)^{1/2}}{(1+a_n^{-1}q^{-n+1})(1+a_n^{-1}q^{-n+2})^{1/2}} \\ &- \frac{q^{-n}\frac{1+a_n^{-1}q}{1+a_n^{-1}q^{-n+1}}}{q^{-n+1/2}\alpha_n^{-1/2}\frac{(1+\alpha_n^{-1}q)^{1/2}(q^{-n}-1)^{1/2}(1-q)^{1/2}}{(1+a_n^{-1}q^{-n+1})(1+a_n^{-1}q^{-n+2})^{1/2}}} \end{aligned} \quad (22)$$

and

$$\frac{b_m^{(n,p,q)}}{\sigma_Y^2} = q^{2n-2} \frac{(1-q^{m-1-n})(1+\alpha_n^{-1}q^{m-1-n})(1+\alpha_n^{-1}q^m)(1-q^m)}{(1+\alpha_n^{-1}q^{2m-2-n})(1+\alpha_n^{-1}q^{2m-1-n})^2(1+\alpha_n^{-1}q^{2m-n})} \cdot \frac{(1+\alpha_n^{-1}q)(q^{-n}-1)(1-q)}{(1+a_n^{-1}q^{-n+1})^2(1+a_n^{-1}q^{-n+2})} \quad (23)$$

Taking the limit $n \rightarrow \infty$ of (22) and (23) and using (7) and (8) , we obtain the recurrence relation of the orthogonal polynomials, say $Q_m(y)$, for the limit distribution, as

$$\begin{aligned} yQ_m(y) &= Q_{m+1}(y) \\ &+ q^{3/2}(1-q)^{-1/2}(q^{-2m-1}(1+q-q^{m+1})-q^{-1})Q_m(y) \\ &+ q^{-4m+4}(1-q)(1-q^m)Q_{m-1}(y), \quad (m \geq 1) \end{aligned} \quad (24)$$

with initial conditions

$$Q_0(y) = 1 \quad Q_1(y) = y.$$

The continuous Stieltjes-Wigert distribution with probability density function

$$v_q^{SW}(x) = \frac{q^{1/8}}{\sqrt{2\pi \log q^{-1}}} \frac{1}{x} e^{\frac{(\log x)^2}{2 \log q}}, \quad x > 0 \quad (25)$$

with

$$\mu = q^{-1}, \quad \sigma^2 = q^{-3}(1-q), \quad (26)$$

is based on the normalized Stieltjes-Wigert orthogonal polynomials, say $p_m^{SW}(x)$, satisfying the recurrence relation

$$\begin{aligned} xp_m^{SW}(x) &= p_{m+1}^{SW}(x) + q^{-2m-1}(1+q-q^{m+1})p_m^{SW}(x) \\ &+ q^{-4m+1}(1-q^m)p_{m-1}^{SW}(x) \quad (m \geq 1) \end{aligned} \quad (27)$$

with initial conditions

$$p_0^{SW}(x) = 1, \quad p_1^{SW}(x) = x - q^{-1}$$

(see Christiansen (2003)).

By standardizing (27) and (25) we obtain respectively (24) and (18) which complete our proof.

ΠΕΡΙΛΗΨΗ

Στην εργασία αυτή, εισάγεται για $0 < q < 1$ μία μετασχηματισμένη q -διωνυμική κατανομή μέσω των κανονικοποιημένων q -Krawtchouk ορθογωνίων πολυωνύμων. Η q -διακριτή αυτή κατανομή ονομάζεται q -διωνυμική κατανομή τύπου I και συμβολίζεται με $\mathcal{B}(n, p, q)$ με $p > 0$. Επίσης, δίδεται μία ερμηνεία αυτής της q -διακριτής κατανομής θεωρώντας, για κατάλληλες τιμές του q , n -διάστατους διανυσματικούς

χώρους επί πεπερασμένων συνόλων με $1/q$ στοιχεία. Επιπλέον, μελετάται η ασυμπτωτική της συμπεριφορά για $n \rightarrow \infty$ αποδεικνύοντας ότι η αντίστοιχη τυποποιημένη q -διωνυμική κατανομή τύπου I συγκλίνει στην τυποποιημένη Stieltjes-Wigert κατανομή.

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